

Sample problems for the first quiz

1. Give a probability space

$$(S = \{\oplus, \ominus\}, \mathcal{F} = 2^S, P = \{p_1, p_2\} \text{ resp. for } S).$$

Note that p_1 is not necessarily equal to p_2 .

(a) Construct a non-deterministic stationary random process $\{\mathbf{x}(t), t \in \mathfrak{R}\}$ that is defined over this probability space.

Hint: Stationarity implies that

$$P\{\zeta \in S : \mathbf{x}(t+c, \zeta) \in A\} = P\{\zeta \in S : \mathbf{x}(t, \zeta) \in A\}$$

for every $A \subset \mathfrak{R}$ and for every $t, c \in \mathfrak{R}$.

(b) Find the mean function and the autocorrelation function of $\mathbf{x}(t)$ that you define in (a).

(c) Verify that the autocorrelation function of $\mathbf{x}(t)$ is non-negative definite.

(d) Construct another non-deterministic stationary random process $\mathbf{y}(t)$ that is different from $\mathbf{x}(t)$ in (a).

Note: Here, “different” means $\mathbf{x}(t, \zeta) \neq \mathbf{y}(t, \zeta)$ for some $\zeta \in S$.

(e) Define a memoryless system $\{\mathbf{T}(x), x \in \mathcal{X}\}$ with system input $\mathbf{x}(t)$ in (a) and system output $\mathbf{y}(t)$ in (d), where $\mathbf{T}(x)$ is also defined over the same probability space.

(f) Is it possible to construct a (weakly) white stationary process $\mathbf{w}(t)$ over the given probability space? Justify your answer.

Note: A process $\mathbf{w}(t)$ is (weakly) *white* if $\mathbf{w}(t_1)$ and $\mathbf{w}(t_2)$ are *uncorrelated* for every $t_1 \neq t_2$.

Solution.

(a) For a stationary process, we shall have

$$P\{\zeta \in S : \mathbf{x}(t+c, \zeta) \in A\} = P\{\zeta \in S : \mathbf{x}(t, \zeta) \in A\}$$

for every $A \in \mathfrak{R}$ and for every $t, c \in \mathfrak{R}$. A non-deterministic process requires $\mathbf{x}(t, \oplus) \neq \mathbf{x}(t, \ominus)$. Thus, we may, for convenience, assign $\mathbf{x}(t, \oplus) = 1$ and $\mathbf{x}(t, \ominus) = -1$. This implies

$$\begin{aligned} P\{\zeta \in S : \mathbf{x}(t+c, \zeta) \in A\} &= P\{\zeta \in S : \mathbf{x}(t, \zeta) \in A\} \\ &= \begin{cases} P\{\} = 0, & \text{if } \pm 1 \notin A; \\ P\{\oplus\} = p_1, & \text{if } 1 \in A \text{ but } -1 \notin A; \\ P\{\ominus\} = p_2, & \text{if } -1 \in A \text{ but } 1 \notin A; \\ P\{\oplus, \ominus\} = 1, & \text{if } \pm 1 \in A. \end{cases} \end{aligned}$$

As a result (of taking $A = \{1\}$ and $A = \{-1\}$), we must have $\mathbf{x}(t+c, \oplus) = 1$ and $\mathbf{x}(t+c, \ominus) = -1$ for every $t, c \in \mathfrak{R}$.

(b)

$$\mu_x(t) = E[\mathbf{x}(t)] = P\{\oplus\} \cdot \mathbf{x}(t, \oplus) + P\{\ominus\} \cdot \mathbf{x}(t, \ominus) = p_1 - p_2$$

and

$$\begin{aligned} R_{xx}(t_1, t_2) &= E[\mathbf{x}(t_1)\mathbf{x}^*(t_2)] = P\{\oplus\} \cdot \mathbf{x}(t_1, \oplus)\mathbf{x}(t_2, \oplus) + P\{\ominus\} \cdot \mathbf{x}(t_1, \ominus)\mathbf{x}(t_2, \ominus) \\ &= p_1 + p_2 = 1. \end{aligned}$$

(c)

$$\sum_i \sum_j a_i a_j^* R_{xx}(t_i, t_j) = \sum_i \sum_j a_i a_j^* = \left(\sum_i a_i \right) \left(\sum_j a_j \right)^* = \left| \sum_i a_i \right|^2 \geq 0.$$

(d) Assign $\mathbf{y}(t, \oplus) = -1$ and $\mathbf{y}(t, \ominus) = 1$.

(e) Since $\mathbf{T}(x)$ is memoryless, it is nothing to do with time argument t . By $\mathbf{T}(\mathbf{x}(t, \zeta), \zeta) = \mathbf{y}(t, \zeta)$, we obtain

$$\mathbf{T}(\mathbf{x}(t, \oplus), \oplus) = \mathbf{y}(t, \oplus) \quad \text{and} \quad \mathbf{T}(\mathbf{x}(t, \ominus), \ominus) = \mathbf{y}(t, \ominus),$$

which implies

$$\mathbf{T}(x, \zeta) = \begin{cases} 1, & x = -1 \text{ and } \zeta = \ominus; \\ -1, & x = 1 \text{ and } \zeta = \oplus; \\ \text{arbitrary,} & x = -1 \text{ and } \zeta = \oplus; \\ \text{arbitrary,} & x = 1 \text{ and } \zeta = \ominus. \end{cases}$$

(f) Let $\mathbf{w}(t, \oplus) = a$ and $\mathbf{w}(t, \ominus) = b$. Then,

$$\mu_w(t) = E[\mathbf{w}(t)] = P\{\oplus\} \cdot \mathbf{x}(t, \oplus) + P\{\ominus\} \cdot \mathbf{x}(t, \ominus) = p_1 a + p_2 b,$$

which implies

$$\begin{aligned} E[\mathbf{w}(t_1)]E[\mathbf{w}^*(t_2)] &= (p_1 a + p_2 b)(p_1 a + p_2 b)^* \\ &= (p_1 a + p_2 b)(p_1 a^* + p_2 b^*) \\ &= p_1^2 |a|^2 + 2p_1 p_2 \text{Re}\{ab^*\} + p_2^2 |b|^2. \end{aligned}$$

In addition,

$$\begin{aligned} R_{ww}(t_1, t_2) &= E[\mathbf{w}(t_1)\mathbf{w}^*(t_2)] \\ &= P\{\oplus\} \cdot \mathbf{w}^*(t_1, \oplus)\mathbf{w}(t_2, \oplus) + P\{\ominus\} \cdot \mathbf{w}^*(t_1, \ominus)\mathbf{w}(t_2, \ominus) \\ &= p_1 |a|^2 + p_2 |b|^2. \end{aligned}$$

Thus, we require

$$\begin{aligned}
p_1^2|a|^2 + 2p_1p_2\text{Re}\{ab^*\} + p_2^2|b|^2 &= p_1|a|^2 + p_2|b|^2 \\
\iff (p_1^2 - p_1)|a|^2 + 2p_1p_2\text{Re}\{ab^*\} + (p_2^2 - p_2)|b|^2 &= 0 \\
\iff (p_1^2 - p_1)|a|^2 - 2(p_1^2 - p_1)\text{Re}\{ab^*\} + (p_1^2 - p_1)|b|^2 &= 0 \\
\iff (p_1^2 - p_1)(|a|^2 - 2\text{Re}\{ab^*\} + |b|^2) &= 0 \\
\iff (p_1^2 - p_1)|a - b|^2 &= 0
\end{aligned}$$

which can be validated for any p_1 (only) under $a = b$. Consequently, $\mathbf{w}(t)$ is a (weakly) white stationary process defined over the given probability space if, and only if, it is a deterministic process.

2. (a) Prove that the autocorrelation function $R_{xx}(t_1, t_2)$ of a random process $\mathbf{x}(t)$ is non-negative definite, namely,

$$\sum_i \sum_j a_i a_j^* R_{xx}(t_i, t_j) \geq 0 \quad \text{for any complex } a_i \text{ and } a_j. \quad (1)$$

- (b) If $\mathbf{x}(t)$ is WSS, determine the equivalent condition to (1) using $S_{xx}(\omega)$, where $S_{xx}(\omega)$ is the power spectrum density of $\mathbf{x}(t)$.

Solution.

- (a)

$$0 \leq E \left[\left| \sum_i a_i \mathbf{x}(t_i) \right|^2 \right] = \sum_i \sum_j a_i a_j^* E[\mathbf{x}(t_i) \mathbf{x}^*(t_j)] = \sum_i \sum_j a_i a_j^* R_{xx}(t_i, t_j).$$

- (b)

$$\begin{aligned}
\sum_i \sum_j a_i a_j^* R_{xx}(t_i, t_j) &\geq 0 \\
\iff \sum_i \sum_j a_i a_j^* R_{xx}(t_i - t_j) &\geq 0 \\
\iff \sum_i \sum_j a_i a_j^* \left(\frac{1}{2\pi} \int_{-\infty}^{\infty} S_{xx}(\omega) e^{j\omega(t_i - t_j)} d\omega \right) &\geq 0 \\
\iff \frac{1}{2\pi} \int_{-\infty}^{\infty} S_{xx}(\omega) \left(\sum_i \sum_j a_i a_j^* e^{j\omega(t_i - t_j)} \right) d\omega &\geq 0 \\
\iff \frac{1}{2\pi} \int_{-\infty}^{\infty} S_{xx}(\omega) \left| \sum_i a_i e^{j\omega t_i} \right|^2 d\omega &\geq 0
\end{aligned}$$

3. A Poisson process $\mathbf{x}(t) \triangleq \mathbf{n}[0, t)$ is one that satisfies

- i) the number of Poisson point occurrences at $\{\mathbf{t}_1, \mathbf{t}_2, \mathbf{t}_3, \dots\}$ in an interval $[t_1, t_2)$ is a Poisson random variable with parameter $\lambda(t_2 - t_1)$, i.e.,

$$\Pr\{\mathbf{n}[t_1, t_2) = k\} = \frac{e^{-\lambda(t_2-t_1)} [\lambda(t_2 - t_1)]^k}{k!},$$

- ii) and $\mathbf{n}[t_1, t_2)$ and $\mathbf{n}[t_3, t_4)$ are independent if $[t_1, t_2)$ and $[t_3, t_4)$ are non-overlapping intervals.

Do the following problems.

- (a) Determine the mean and autocorrelation functions of $\mathbf{x}(t)$.
 (b) Determine $E[\mathbf{x}(t_1)\mathbf{x}(t_2)\mathbf{x}(t_3)]$ for $t_1 < t_2 < t_3$.

Note: $E[\mathbf{n}^3[0, t)] = \lambda^3 t^3 + 3\lambda^2 t^2 + \lambda t$ and $E[\mathbf{n}^2[0, t)] = \lambda^2 t^2 + \lambda t$.

Solution.

- (a) See Slides 9-48 and 9-49.
 (b) For $t_1 < t_2 < t_3$,

$$\begin{aligned} E[\mathbf{x}(t_1)\mathbf{x}(t_2)\mathbf{x}(t_3)] &= E[\mathbf{n}[0, t_1) \cdot (\mathbf{n}[0, t_1) + \mathbf{n}[t_1, t_2)) \cdot (\mathbf{n}[0, t_1) + \mathbf{n}[t_1, t_2) + \mathbf{n}[t_2, t_3))] \\ &= E[\mathbf{n}^3[0, t_1)] + E[\mathbf{n}^2[0, t_1)\mathbf{n}[t_1, t_2)] + E[\mathbf{n}^2[0, t_1)\mathbf{n}[t_2, t_3)] \\ &\quad + E[\mathbf{n}^2[0, t_1)\mathbf{n}[t_1, t_2)] + E[\mathbf{n}[0, t_1)\mathbf{n}^2[t_1, t_2)] + E[\mathbf{n}[0, t_1)\mathbf{n}[t_1, t_2)\mathbf{n}[t_2, t_3)] \\ &= E[\mathbf{n}^3[0, t_1)] + E[\mathbf{n}^2[0, t_1)]E[\mathbf{n}[t_1, t_2)] + E[\mathbf{n}^2[0, t_1)]E[\mathbf{n}[t_1, t_3)] \\ &\quad + E[\mathbf{n}[0, t_1)]E[\mathbf{n}^2[t_1, t_2)] + E[\mathbf{n}[0, t_1)]E[\mathbf{n}[t_1, t_2)]E[\mathbf{n}[t_2, t_3)] \\ &= (\lambda^3 t_1^3 + 3\lambda^2 t_1^2 + \lambda t_1) + (\lambda^2 t_1^2 + \lambda t_1)(\lambda t_2 - \lambda t_1) + (\lambda^2 t_1^2 + \lambda t_1)(\lambda t_3 - \lambda t_1) \\ &\quad + \lambda t_1(\lambda^2(t_2 - t_1)^2 + \lambda(t_2 - t_1)) + \lambda t_1(\lambda t_2 - \lambda t_1)(\lambda t_3 - \lambda t_2) \\ &= \lambda^3 t_1^3 + 3\lambda^2 t_1^2 + \lambda t_1 + \lambda^3 t_1^2 t_2 - \lambda^3 t_1^3 + \lambda^2 t_1 t_2 - \lambda^2 t_1^2 + \lambda^3 t_1^2 t_3 - \lambda^3 t_1^3 + \lambda^2 t_1 t_3 - \lambda^2 t_1^2 \\ &\quad + \lambda^3 t_1^3 + \lambda^3 t_1 t_2^2 - 2\lambda^3 t_1^2 t_2 + \lambda^2 t_1 t_2 - \lambda^2 t_1^2 + \lambda^3 t_1 t_2 t_3 - \lambda^3 t_1 t_2^2 - \lambda^3 t_1^2 t_3 + \lambda^3 t_1^2 t_2 \\ &= \lambda t_1 + \lambda^2(2t_1 t_2 + t_1 t_3) + \lambda^3 t_1 t_2 t_3 \\ &\quad \text{or} = \lambda^2 t_1 t_2 + \lambda t_1(1 + \lambda t_2)(1 + \lambda t_3). \end{aligned}$$

4. (a) Prove that the process $\mathbf{x}(t) = \mathbf{a} \cos(\omega t) + \mathbf{b} \sin(\omega t)$ is WSS if, and only if, \mathbf{a} and \mathbf{b} are uncorrelated zero mean with equal variance.
 (b) Disprove the statement that if the process $\mathbf{x}(t) = a \cos(\omega t + \varphi)$ is WSS, then φ is uniformly distributed over $[-\pi, \pi)$.
 (c) Disprove the statement that if the complex process $\mathbf{z}(t) = a e^{j(\omega t + \varphi)}$ is WSS, then φ is uniformly distributed over $[-\pi, \pi)$.

Solution.

- (a) The “if” proof has been provided in my slides. It remains to prove the “only if” part.

WSS of $\mathbf{x}(t) = \mathbf{a} \cos(\omega t) + \mathbf{b} \sin(\omega t)$ implies that

$$E[\mathbf{x}(t)] = E[\mathbf{a}] \cos(\omega t) + E[\mathbf{b}] \sin(\omega t) = C \text{ for any } t$$

which in turn implies that $E[\mathbf{a}] = C$ at $t = 0$ and $E[\mathbf{a}] = -C$ at $t = \pi/\omega$. Hence, $E[\mathbf{a}] = C = 0$. As a result, we have $E[\mathbf{b}] \sin(\omega t) = 0$, which immediately gives $E[\mathbf{b}] = 0$.

Next, we derive

$$\begin{aligned} E[\mathbf{x}(t+\tau)\mathbf{x}(t)] &= E[\mathbf{a}^2] \cos(\omega t) \cos(\omega(t+\tau)) + E[\mathbf{b}^2] \sin(\omega t) \sin(\omega(t+\tau)) \\ &\quad + E[\mathbf{a}\mathbf{b}] \sin(\omega(2t+\tau)) \\ &= E[\mathbf{a}^2] \frac{\cos(\omega\tau) + \cos(\omega(2t+\tau))}{2} + E[\mathbf{b}^2] \frac{\cos(\omega\tau) - \cos(\omega(2t+\tau))}{2} \\ &\quad + E[\mathbf{a}\mathbf{b}] \sin(\omega(2t+\tau)) \\ &= \frac{E[\mathbf{a}^2] + E[\mathbf{b}^2]}{2} \cos(\omega\tau) + \frac{E[\mathbf{a}^2] - E[\mathbf{b}^2]}{2} \cos(\omega(2t+\tau)) \\ &\quad + E[\mathbf{a}\mathbf{b}] \sin(\omega(2t+\tau)) \end{aligned}$$

which is only a function of τ . Hence, $E[\mathbf{a}^2] = E[\mathbf{b}^2]$ and $E[\mathbf{a}\mathbf{b}] = 0$.

- (b) To disprove the statement, it suffices to provide a counterexample. Suppose $\varphi = 0, \pi/2, \pi$ and $3\pi/2$ with equal probability and is independent of ω . Then,

$$\begin{aligned} E[\mathbf{x}(t)] &= aE[\cos(\omega t + \varphi)] \\ &= a\frac{1}{4}E[\cos(\omega t)] - a\frac{1}{4}E[\sin(\omega t)] - a\frac{1}{4}E[\cos(\omega t)] + a\frac{1}{4}E[\sin(\omega t)] \\ &= 0 \end{aligned}$$

and

$$\begin{aligned} E[\mathbf{x}(t_1)\mathbf{x}(t_2)] &= E\{a^2 \cos(\omega t_1 + \varphi) \cos(\omega t_2 + \varphi)\} \\ &= E\left\{a^2 \frac{\cos[\omega(t_1 - t_2)] + \cos[\omega(t_1 + t_2) + 2\varphi]}{2}\right\} = \frac{a^2}{2}E[\cos(\omega(t_1 - t_2))]. \end{aligned}$$

- (c) Again, a counterexample is provided. Suppose $\varphi = 0$ and π with equal probability and is independent of ω . Then,

$$\begin{aligned} E[\mathbf{z}(t)] &= E[ae^{j(\omega t + \varphi)}] \\ &= aE[e^{j\omega t}]E[e^{j\varphi}] = 0 \end{aligned}$$

and

$$\begin{aligned} E[\mathbf{z}(t)\mathbf{z}^*(t+\tau)] &= E[ae^{j(\omega t + \varphi)}a^*e^{-j(\omega(t+\tau) + \varphi)}] \\ &= |a|^2E[e^{j\omega\tau}] \end{aligned}$$